

Association between maternal age at childbirth and child and adult outcomes in the offspring: a prospective study in five low-income and middle-income countries (COHORTS collaboration)



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Summary

Background Both young and advanced maternal age is associated with adverse birth and child outcomes. Few studies have examined these associations in low-income and middle-income countries (LMICs) and none have studied adult outcomes in the offspring. We aimed to examine both child and adult outcomes in five LMICs.

Methods In this prospective study, we pooled data from COHORTS (Consortium for Health Orientated Research in Transitioning Societies)—a collaboration of five birth cohorts from LMICs (Brazil, Guatemala, India, the Philippines, and South Africa), in which mothers were recruited before or during pregnancy, and the children followed up to adulthood. We examined associations between maternal age and offspring birthweight, gestational age at birth, height-for-age and weight-for-height Z scores in childhood, attained schooling, and adult height, body composition (body-mass index, waist circumference, fat, and lean mass), and cardiometabolic risk factors (blood pressure and fasting plasma glucose concentration), along with binary variables derived from these. Analyses were unadjusted and adjusted for maternal socioeconomic status, height and parity, and breastfeeding duration.

Findings We obtained data for 22 188 mothers from the five cohorts, enrolment into which took place at various times between 1969 and 1989. Data for maternal age and at least one outcome were available for 19 403 offspring (87%). In unadjusted analyses, younger (≤ 19 years) and older (≥ 35 years) maternal age were associated with lower birthweight, gestational age, child nutritional status, and schooling. After adjustment, associations with younger maternal age remained for low birthweight (odds ratio [OR] 1.18 [95% CI 1.02–1.36]), preterm birth (1.26 [1.03–1.53]), 2-year stunting (1.46 [1.25–1.70]), and failure to complete secondary schooling (1.38 [1.18–1.62]) compared with mothers aged 20–24 years. After adjustment, older maternal age remained associated with increased risk of preterm birth (OR 1.33 [95% CI 1.05–1.67]), but children of older mothers had less 2-year stunting (0.64 [0.54–0.77]) and failure to complete secondary schooling (0.59 [0.48–0.71]) than did those with mothers aged 20–24 years. Offspring of both younger and older mothers had higher adult fasting glucose concentrations (roughly 0.05 mmol/L).

Interpretation Children of young mothers in LMICs are disadvantaged at birth and in childhood nutrition and schooling. Efforts to prevent early childbearing should be strengthened. After adjustment for confounders, children of older mothers have advantages in nutritional status and schooling. Extremes of maternal age could be associated with disturbed offspring glucose metabolism.

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Introduction

Young maternal age at childbearing (≤ 19 years) is associated with an increased risk of preterm birth and intrauterine growth restriction, infant mortality, and child undernutrition.^{1–12} These associations result from behavioural, social, and biological factors. Younger mothers might breastfeed for a shorter duration than older mothers^{3,13} and be behaviourally immature and therefore less able to attend to their infant's needs. They tend to have lower socioeconomic status, less schooling, and less stable partnerships than older mothers. If still growing, their nutritional needs compete with those of the fetus.¹⁴

Advanced maternal age (≥ 35 years) is associated with increased stillbirths, preterm births, intrauterine growth restriction, as well as young maternal age, and chromosomal abnormalities.^{15–18} Again, these consequences result from multiple factors. In some settings, older mothers have lower socioeconomic status, less schooling, and higher parity, whereas in others they are educated women who have delayed pregnancy for career reasons. Older mothers are at increased risk of obesity, diabetes, hypertension, and associated pregnancy complications.

Recent interest has grown in the developmental origins of human capital and adult disease, especially

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See [Comment](#) page e342

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See [Online](#) for appendix

cardiovascular disease and its risk factors hypertension and diabetes.¹⁹ Studies from high-income countries show lower educational achievement in children of younger mothers.²⁰ Some studies have further shown increased childhood blood pressure in the offspring of older mothers.^{21–24} No published studies have examined associations between maternal age and adult risk factors for cardiovascular disease.

Most studies linking maternal age to child outcomes are from high-income countries (HICs). However, young maternal age is more common in low-income and middle-income countries (LMICs; fertility rate among women aged younger than 19 years roughly 103 per 1000 women per year) than in HICs (21 per 1000).²⁵ More young mothers in LMICs could be physically immature because of delayed completion of growth among undernourished girls. Confounding effects of socioeconomic status might differ in LMICs—for example, in some settings, young age at marriage is a norm and not necessarily associated with deprivation.^{3,10}

We aimed to explore associations between maternal age and birth outcomes, child nutritional status and schooling, and adult size, body composition, and cardiometabolic risk factors. We adjusted for socioeconomic confounding factors, and explored maternal height, breastfeeding duration, and parity as potential mediators of associations.

Methods

Study design

COHORTS (Consortium for Health Orientated Research in Transitioning Societies) is a collaboration of five birth cohorts from LMICs, in which mothers were recruited before or during pregnancy, and the children followed up to adulthood.²⁶ In this prospective study, the cohorts include the 1982 Pelotas Cohort (Brazil); the Institute of Nutrition of Central America and Panama Nutrition Trial Cohort (Guatemala); the New Delhi Cohort (India); the Cebu Longitudinal Health and Nutrition Survey (Philippines); and the Birth to Twenty Cohort (South Africa; appendix p 1).²⁶ All studies were approved by appropriate institutional ethics committees. Informed verbal or written consent was obtained at recruitment from the mothers in the original birth cohort studies. Informed verbal or written consent was obtained at each round of follow-up, from a parent for childhood follow-ups and from the cohort member themselves for the adult follow-up.

Procedures

The mother's age at the birth of the index child was calculated from data obtained at interview before or during pregnancy, or, for South African data, from birth notification forms.

Birthweight was measured by researchers (Brazil, India, and Guatemala) or obtained from hospital records (the Philippines and South Africa). Gestational age was

calculated from the last menstrual period date obtained by prospective surveillance (Guatemala and India), from the mother at recruitment (the Philippines), or from medical records (Brazil and South Africa). In the Philippines, gestational age was obtained for low birthweight babies by newborn clinical assessment.²⁷ Low birthweight was defined as <2500 g, preterm birth as gestational age <37 weeks, and smallness-for-gestational-age as birthweight below the age-specific and sex-specific 10th percentile of a US reference population.²⁸

In all sites, post-natal weight and height were measured longitudinally with standardised methods. Measurements at 2 years were available in all sites, and 4-year measurements in all except the Philippines, where the next available age (8·5 years) was used to define so-called mid-childhood size. Height-for-age and weight-for-height were converted into Z-scores (HAZ and WHZ, respectively) with the WHO growth reference. Stunting and wasting were defined as HAZ and WHZ below –2 SDs, respectively. Breastfeeding data were recorded prospectively with different methods in each cohort; for this analysis we used duration in months of any breastfeeding (as opposed to exclusive breastfeeding), which was available for all cohorts except India.

Height, weight, and waist circumference were measured with standardised techniques. Fat and fat-free mass were measured with site-specific methods, as previously described.²⁹ In Brazil, bioelectrical impedance was measured and results corrected based on a validation study that used isotopic methods; data are only available for men. In Guatemala, weight, height, and abdominal circumference were measured and entered into a hydrostatic-weighing validated equation. The Indian and Filipino cohorts used published equations that have been validated for use in Asian populations for estimation of body fat from skinfold measures. South Africa used dual X-ray absorptiometry (Hologic Delphi, Bedford, MA, USA). Fat mass was calculated as: % body fat × body weight, and fat-free mass as weight – fat mass.

Overweight was defined as a body-mass index (BMI) of 25 kg/m² or more and obesity as BMI 30 kg/m² or more. Blood pressure was measured seated, after a 5–10 min rest, with appropriate cuff sizes and a variety of devices, as previously described.³⁰ High blood pressure was defined as systolic blood pressure 130 mm Hg or greater or diastolic blood pressure 85 mm Hg or greater.³¹ Fasting glucose was measured in all sites except Brazil, where a random sample was collected and glucose values adjusted for time since the last meal.³² Impaired fasting glucose was defined as a fasting glucose concentration of 6·1 mmol/L or greater but less than 7·0 mmol/L and diabetes as a concentration of 7·0 mmol/L or greater.³³ Pregnant women were excluded from all these analyses.

Socioeconomic status was considered a potential confounding factor and was assessed with five variables: maternal schooling, marital status, wealth index, urban

For the WHO growth reference data see www.who.int/growthref/en

| | Brazil | | Guatemala | | India | | Philippines | | South Africa | |
|--|--------------------|------|--------------------|------|--------------------|------|--------------------|------|--------------------|------|
| | Mean (SD) or n (%) | N | Mean (SD) or n (%) | N | Mean (SD) or n (%) | N | Mean (SD) or n (%) | N | Mean (SD) or n (%) | N |
| Original cohort | .. | 5913 | .. | 2392 | .. | 7530 | .. | 3080 | .. | 3273 |
| Maternal age | .. | 5912 | .. | 2344 | .. | 5395 | .. | 3080 | .. | 3271 |
| Maternal age and any outcome of interest | .. | 5910 | .. | 1942 | .. | 5203 | .. | 3078 | .. | 3270 |
| Maternal age (years) | 25.8 (6.1) | 5910 | 27.2 (7.2) | 1942 | 26.0 (5.2) | 5203 | 26.3 (6.0) | 3078 | 26.0 (6.1) | 3270 |
| Child outcomes | | | | | | | | | | |
| Birthweight (g) | 3190 (568) | 5804 | 3048 (492) | 973 | 2792 (440) | 4832 | 2988 (438) | 3029 | 3070 (513) | 3265 |
| Low birthweight | 640 (11.0%) | 5804 | 136 (14.0%) | 973 | 1393 (28.8%) | 4832 | 448 (14.8%) | 3029 | 420 (12.9%) | 3265 |
| Small for gestational age | 683 (14.9%) | 4596 | 233 (31.1%) | 748 | 361 (40.9%) | 883 | 735 (24.5%) | 2999 | 453 (14.3%) | 3166 |
| Gestation (weeks) | 39.3 (1.9) | 4600 | 39.3 (3.0) | 851 | 38.9 (2.6) | 937 | 38.7 (2.2) | 3048 | 38.1 (1.9) | 3169 |
| Preterm birth | 289 (6.3%) | 4600 | 109 (12.8%) | 851 | 141 (15.0%) | 937 | 466 (15.3%) | 3048 | 388 (12.2%) | 3169 |
| Height (Z) at 2 years | -0.67 (1.25) | 4835 | -2.91 (1.15) | 1079 | -2.01 (1.23) | 3697 | -2.41 (1.15) | 2504 | -1.16 (1.12) | 1804 |
| Weight-for-height (Z) at 2 years | 0.56 (0.99) | 4832 | -0.28 (0.90) | 1076 | -0.66 (1.03) | 3666 | -0.62 (0.96) | 2502 | 0.19 (1.35) | 1803 |
| Stunted at 2 years | 667 (13.8%) | 4835 | 858 (79.5%) | 1079 | 1841 (49.8%) | 3697 | 1558 (62.2%) | 2504 | 372 (20.6%) | 1804 |
| Wasted at 2 years | 40 (0.8%) | 4832 | 52 (4.8%) | 1076 | 362 (9.9%) | 3666 | 182 (7.3%) | 2502 | 86 (4.8%) | 1803 |
| Height (Z) MC | -0.66 (1.11) | 4651 | -2.42 (0.97) | 998 | -1.98 (1.06) | 3197 | -2.05 (0.94) | 2261 | -0.91 (0.93) | 1804 |
| Weight-for-height (Z) MC | 0.63 (1.01) | 4648 | 0.32 (0.83) | 998 | -0.17 (0.88) | 3173 | -0.81 (0.89) | 2261 | 0.18 (1.03) | 1804 |
| Stunted mid-childhood | 505 (10.9%) | 4651 | 672 (67.3%) | 998 | 1547 (48.4%) | 3197 | 1215 (53.7%) | 2261 | 212 (11.8%) | 1804 |
| Wasted mid-childhood | 15 (0.3%) | 4648 | 2 (0.2%) | 998 | 72 (2.3%) | 3173 | 195 (8.6%) | 2261 | 29 (1.6%) | 1804 |
| Adult outcomes | | | | | | | | | | |
| Age (years) | 22.7 (0.4) | 4193 | 32.4 (4.1) | 1297 | 29.1 (1.3) | 1038 | 21.2 (0.9) | 2040 | 18.1 (0.5) | 1989 |
| Attained schooling (years) | 9.3 (3.2) | 4076 | 4.7 (3.4) | 1458 | 13.2 (3.4) | 1038 | 10.0 (3.0) | 2069 | 11.1 (1.5) | 2036 |
| Failed to finish secondary school | 1974 (48.4%) | 4076 | 1369 (93.9%) | 1458 | 145 (14.0%) | 1038 | 613 (29.6%) | 2069 | 781 (38.4%) | 2036 |
| Height (cm) | | | | | | | | | | |
| Men | 173.7 (6.9) | 2207 | 162.8 (6.0) | 552 | 169.3 (6.4) | 590 | 163.0 (5.9) | 1079 | 171.0 (7.5) | 933 |
| Women | 160.8 (6.2) | 1979 | 150.6 (5.6) | 608 | 154.3 (5.1) | 446 | 151.2 (5.5) | 958 | 159.7 (6.3) | 1000 |
| Short stature | 167 (4.0%) | 4186 | 535 (46.1%) | 1160 | 161 (15.5%) | 1036 | 864 (42.4%) | 2037 | 151 (7.8%) | 1933 |
| Body-mass index (kg/m ²) | 23.6 (4.4) | 4185 | 25.9 (4.4) | 1160 | 24.4 (4.6) | 1036 | 20.6 (3.1) | 2032 | 21.7 (4.2) | 1932 |
| Waist circumference (cm) | | | | | | | | | | |
| Men | 80.9 (10.1) | 2205 | 86.7 (9.1) | 540 | 88.9 (11.7) | 590 | 72.0 (7.5) | 1077 | .. | .. |
| Women | 74.9 (10.6) | 1978 | 92.4 (12.1) | 604 | 78.6 (12.4) | 448 | 67.6 (7.2) | 951 | .. | .. |
| Overweight | 1190 (28.4%) | 4185 | 603 (52.0%) | 1160 | 465 (44.9%) | 1036 | 176 (8.7%) | 2032 | 330 (17.1%) | 1932 |
| Obese | 341 (8.1%) | 4185 | 202 (17.4%) | 1160 | 94 (9.1%) | 1036 | 32 (1.6%) | 2032 | 104 (5.4%) | 1932 |
| Fat mass (kg) | | | | | | | | | | |
| Men | 12.2 (5.1) | 1952 | 14.0 (6.6) | 540 | 17.2 (6.9) | 588 | 9.7 (4.5) | 988 | 8.7 (5.0) | 879 |
| Women | .. | .. | 22.3 (9.1) | 602 | 20.3 (8.1) | 446 | 15.4 (4.6) | 822 | 19.4 (7.7) | 936 |
| Lean mass (kg) | | | | | | | | | | |
| Men | 60.3 (9.7) | 1952 | 51.4 (5.4) | 540 | 53.4 (7.8) | 588 | 46.4 (5.8) | 988 | 51.0 (6.7) | 879 |
| Women | .. | .. | 38.9 (3.6) | 602 | 37.6 (5.6) | 446 | 31.0 (4.2) | 822 | 39.0 (5.4) | 936 |
| Body fat (%) | | | | | | | | | | |
| Men | 16.3 (3.8) | 2196 | 20.5 (6.6) | 540 | 23.5 (6.0) | 588 | 16.7 (5.1) | 988 | 14.1 (5.3) | 883 |
| Women | .. | .. | 35.1 (7.3) | 602 | 33.7 (7.5) | 446 | 32.7 (4.8) | 822 | 32.3 (6.6) | 939 |
| Systolic blood pressure (mm Hg) | 117.7 (15.0) | 4188 | 112.1 (12.9) | 1248 | 113.0 (12.2) | 1033 | 105.9 (12.2) | 2036 | 117.6 (10.7) | 1871 |
| Diastolic blood pressure (mm Hg) | 73.8 (11.4) | 4188 | 70.8 (9.5) | 1248 | 75.3 (10.0) | 1033 | 72.1 (10.0) | 2042 | 71.4 (8.5) | 1871 |
| High blood pressure | 1384 (33.0%) | 4188 | 224 (17.9%) | 1248 | 342 (33.1%) | 1033 | 589 (28.8%) | 2044 | 400 (21.4%) | 1871 |
| Fasting glucose (mmol/L) | 5.02 (0.73) | 3612 | 5.20 (1.11) | 985 | 5.43 (1.01) | 1013 | 4.63 (0.55) | 1703 | 4.62 (0.45) | 1189 |
| IFG or diabetes | 281 (7.8%) | 3618 | 60 (6.1%) | 991 | 184 (18.1%) | 1014 | 16 (0.9%) | 1703 | 10 (0.8%) | 1194 |

(Table 1 continues on next page)

| | Brazil | | Guatemala | | India | | Philippines | | South Africa | |
|--------------------------------|--------------------|------|--------------------|------|--------------------|------|--------------------|------|--------------------|------|
| | Mean (SD) or n (%) | N | Mean (SD) or n (%) | N | Mean (SD) or n (%) | N | Mean (SD) or n (%) | N | Mean (SD) or n (%) | N |
| (Continued from previous page) | | | | | | | | | | |
| Covariates | | | | | | | | | | |
| Mother married | 5422 (91.8%) | 5905 | 1557 (91.5%) | 1701 | 5200 (100.0%) | 5201 | 2963 (97.5%) | 3038 | 1414 (43.5%) | 3248 |
| Maternal education (years) | 6.5 (4.2) | 5910 | 1.3 (1.6) | 1913 | 5.1 (4.6) | 5203 | 7.1 (3.3) | 3078 | 9.5 (3.0) | 3048 |
| Maternal wealth | 3.3 (1.2) | 4988 | -0.12 (0.93) | 1784 | 4.1 (1.2) | 5190 | 3.5 (2.6) | 3077 | 3.9 (1.6) | 2857 |
| Ethnic origin | | | | | | | | | | |
| White | 3457 (75.5%) | 4580 | .. | 1942 | .. | 5203 | .. | 3078 | 207 (6.3%) | 3270 |
| Black | 711 (15.5%) | 4580 | .. | 1942 | .. | 5203 | .. | 3078 | 2567 (78.5%) | 3270 |
| Asian | 76 (1.7%) | 4580 | .. | 1942 | 5203 (100%) | 5203 | 3078 (100%) | 3078 | 115 (3.5%) | 3270 |
| Other | 336 (7.3%) | 4580 | 1942 (100%) | 1942 | .. | 5203 | .. | 3078 | 381 (11.7%) | 3270 |
| Urban score | Urban | .. | Rural | .. | Urban | .. | 30.6 (12.6) | 3078 | Urban | .. |
| Maternal height (cm) | 156.4 (6.0) | 5804 | 148.5 (5.2) | 1604 | 151.8 (5.5) | 1296 | 150.6 (5.0) | 3078 | 158.8 (6.1) | 1757 |
| Breastfeeding | | | | | | | | | | |
| None | 420 (9.4%) | 4489 | 7 (0.7%) | 988 | .. | .. | 204 (7.6%) | 2668 | 203 (8.5%) | 2390 |
| <6 months | 2562 (57.1%) | 4489 | 40 (4.0%) | 988 | .. | .. | 397 (14.9%) | 2668 | 745 (31.2%) | 2390 |
| <1 year | 690 (15.4%) | 4489 | 149 (15.1%) | 988 | .. | .. | 353 (13.2%) | 2668 | 368 (15.4%) | 2390 |
| ≥1 year | 817 (18.2%) | 4489 | 792 (80.2%) | 988 | .. | .. | 1714 (64.2%) | 2668 | 1074 (44.9%) | 2390 |
| Maternal parity | | | | | | | | | | |
| 1 | 2321 (39.3%) | 5908 | 310 (16.0%) | 1938 | 77.2 (17.2%) | 4492 | 686 (22.3%) | 3078 | 1197 (36.6%) | 3270 |
| 2 | 1659 (28.1%) | 5908 | 295 (15.2%) | 1938 | 1140 (25.4%) | 4492 | 693 (22.5%) | 3078 | 988 (30.2%) | 3270 |
| 3 | 965 (16.3%) | 5908 | 272 (14.0%) | 1938 | 1000 (22.3%) | 4492 | 598 (19.4%) | 3078 | 578 (17.7%) | 3270 |
| 4+ | 963 (16.3%) | 5908 | 1061 (54.7%) | 1938 | 1580 (35.2%) | 4492 | 1101 (35.8%) | 3078 | 507 (15.5%) | 3270 |

Breastfeeding duration was not available for India. Imputed values were used for maternal schooling (Brazil, n=7 [0.1%]; Guatemala, n=64 [3.3%]; India, n=9 [0.17%]; Philippines, n=0; South Africa, n=341 [10.4%]) and maternal wealth (Brazil, n=922 [15.6%]; Guatemala, n=158 [8.1%]; India, n=13 [2.5%]; Philippines, n=1 [0.03%]; South Africa, n=413 [12.6%]). Maternal height was missing for 106 in Brazil (1.8%), 338 in Guatemala (17.4%), 3907 in India (75.1%), 0 in the Philippines, and 1513 (46.3%) in South Africa. IFG=impaired fasting glucose. MC=mid-childhood.

Table 1: Maternal age, birth, child, and adult outcomes, socioeconomic variables and maternal height, breastfeeding, and parity in the five cohorts

or rural residence, and ethnic origin. Wealth index was a score derived in each cohort based on type of housing and ownership of household assets (appendix p 2). The Brazilian, Indian, and South African cohorts are urban, and the Guatemalan cohort is rural; in the mixed Filipino cohort, a so-called urbanicity index was used.³⁴ Brazil and South Africa had white, black, Asian, and other ethnic subgroups. Maternal height was potentially both a confounder of maternal age effects (due to secular trends in height) and a mediator (due to younger mothers not having attained final height). Maternal parity and breastfeeding duration were potential mediators (older mothers tend to have higher parity and younger mothers might breastfeed for a shorter time). Parity was coded as 1, 2, 3, or 4 or more.

Statistical analysis

The number of offspring for whom outcomes were available diminished with age at follow-up; for example, n=17 903 (81%) for birthweight and n=10 376 (47%) for adult blood pressure. The maternal age distribution was similar in successive waves of follow-up (appendix p 3). For each outcome, we used the maximum sample with available data. To test the representativeness of our analysis sample, we compared maternal age in those

included in the analysis (with data for any childhood or adult outcome of interest) and those not included, using *t* tests (appendix p 4). We used maternal age as a continuous variable where possible, but used categories (≤19 years, ≥35 years, and 5-year intervening bands) to check for linearity and for tables, figures, and odds ratio calculations. Percentage of body fat and wealth index were non-normally distributed and were Fisher-Yates transformed.³⁵

We first analysed associations between maternal age and outcomes in each cohort with multiple linear regression for continuous outcomes and multiple logistic regression for dichotomous outcomes. We assessed non-linear associations with quadratic terms. We then produced pooled analyses, including main effects for each cohort and interaction terms for cohort and control variables. We tested for heterogeneity among cohorts with *F* tests, comparing sums of squares explained when effects were or were not allowed to vary across sites.³⁵ We used a sequence of regression models: (1) adjusted for sex, and adult age (adult outcomes only); (2) further adjusted for socioeconomic variables; (3) further adjusted for maternal height; (4) further adjusted for breastfeeding duration; and (5) further adjusted for parity. Missing

| | Maternal age (years) | Birthweight (g) | Gestational age (weeks) | Height at 2 years (Z) | Weight-for-height at age 2 years (Z) | Schooling (years) | Adult height (cm) | Systolic blood pressure (mm Hg) | Fasting glucose (mmol/L) |
|---|----------------------|-----------------|-------------------------|-----------------------|--------------------------------------|-------------------|-------------------|---------------------------------|--------------------------|
| Maternal schooling (correlation) | | | | | | | | | |
| Brazil | 0.06‡ | 0.09‡ | -0.03 | 0.31‡ | 0.14‡ | 0.51‡ | 0.12‡ | -0.01 | -0.03 |
| Guatemala | -0.05* | 0.08* | -0.01 | 0.06 | 0.01 | 0.28‡ | 0.12‡ | 0 | -0.06* |
| India | -0.12‡ | 0.17‡ | -0.02 | 0.45‡ | 0.28‡ | 0.38‡ | 0.12‡ | 0 | 0.02 |
| Philippines | -0.05† | 0.05* | 0.04 | 0.33‡ | 0.13‡ | 0.46‡ | 0.14‡ | 0.04 | -0.03 |
| South Africa | -0.15‡ | 0.06† | 0.06† | 0.16‡ | -0.01 | 0.21‡ | 0.05* | -0.09‡ | -0.03 |
| Marital status (mean values) | | | | | | | | | |
| Brazil | | | | | | | | | |
| Married | 26.1 | 3203 | 39.3 | -0.65 | 0.56 | 9.4 | 167.6 | 117.8 | 5.03 |
| Single | 22.5‡ | 3046‡ | 39.3 | -0.91‡ | 0.53 | 8.4‡ | 167.1 | 116.5 | 4.99 |
| Guatemala | | | | | | | | | |
| Married | 27.9 | 3075 | 39.4 | -2.88 | -0.26 | 4.7 | 156.6 | 112.2 | 5.23 |
| Single | 26.1† | 2896† | 38.4* | -3.20* | -0.53† | 4.0* | 155.8 | 111.7 | 5.18 |
| Philippines | | | | | | | | | |
| Married | 26.5 | 2993 | 38.7 | -2.41 | -0.62 | 10.0 | 157.5 | 105.8 | 4.63 |
| Single | 21.8‡ | 2852† | 38.5 | -2.34 | -0.58 | 9.9 | 157.6 | 105.3 | 4.45* |
| South Africa | | | | | | | | | |
| Married | 28.4 | 3098 | 38.4 | -1.02 | 0.02 | 11.1 | 166.1‡ | 118.1 | 4.64 |
| Single | 24.1‡ | 3051† | 38.0‡ | -1.25‡ | 0.30‡ | 11.0 | 164.6 | 117.3 | 4.61 |
| Wealth (correlation) | | | | | | | | | |
| Brazil | 0.22‡ | 0.11‡ | -0.04* | 0.34‡ | 0.17‡ | 0.47‡ | 0.10‡ | 0 | -0.06† |
| Guatemala | 0.02 | 0.11† | -0.01 | 0.14‡ | 0.09† | 0.23‡ | 0.10† | 0.03 | -0.01 |
| India | 0.11‡ | 0.13‡ | -0.01 | 0.23‡ | 0.17‡ | 0.22‡ | 0.05 | 0.03 | 0.02 |
| Philippines | 0.09‡ | 0.08‡ | 0.02 | 0.26‡ | 0.15‡ | 0.24‡ | 0.12‡ | 0.04 | -0.01 |
| South Africa | 0.02 | 0.02 | 0.06† | 0.19‡ | -0.05 | 0.16‡ | 0.03 | -0.06* | 0.03 |
| Urban (correlation) | | | | | | | | | |
| Philippines | -0.06† | 0.04* | -0.01 | 0.15‡ | -0.04 | 0.20‡ | 0.04 | 0.01 | -0.02 |
| Ethnic origin (mean values) | | | | | | | | | |
| Brazil | | | | | | | | | |
| White | 26.0 | 3247‡ | 39.4 | -0.55‡ | 0.61‡ | 9.8‡ | 167.9* | 117.3† | 5.02 |
| Black | 25.8 | 3139‡ | 39.2 | -1.10‡ | 0.44‡ | 8.2‡ | 166.7* | 119.4† | 5.00 |
| Asian | 25.2 | 3225‡ | 39.7 | -0.98‡ | 0.38‡ | 8.3‡ | 166.1* | 115.8† | 5.04 |
| Other | 26.3 | 3207‡ | 39.3 | -0.99‡ | 0.47‡ | 7.8‡ | 167.1* | 119.1† | 5.13 |
| South Africa | | | | | | | | | |
| White | 27.8‡ | 3197‡ | 39.0‡ | -0.37‡ | 0.36‡ | 11.6‡ | 169.8† | 114.3‡ | NA |
| Black | 25.9‡ | 3077‡ | 37.9‡ | -1.24‡ | 0.35‡ | 11.1‡ | 165.0† | 118.0‡ | 4.62 |
| Asian | 26.3‡ | 2911‡ | 38.8‡ | -0.77‡ | -0.58‡ | 11.6‡ | 167.1† | 111.0‡ | 4.86 |
| Other | 25.6‡ | 3007‡ | 38.9‡ | -1.13‡ | -0.51‡ | 10.7‡ | 164.9† | 116.1‡ | 4.63 |
| Maternal height (correlation) | | | | | | | | | |
| Brazil | 0 | 0.18‡ | 0.04† | 0.38‡ | 0.13‡ | 0.19‡ | 0.38‡ | 0.03 | -0.01 |
| Guatemala | -0.01 | 0.14‡ | -0.02 | 0.30‡ | 0.07* | 0.10‡ | 0.32‡ | 0.03 | 0.01 |
| India | -0.02 | 0.09‡ | 0 | 0.20‡ | 0.08‡ | 0.04 | 0.14‡ | -0.03 | 0.07* |
| Philippines | 0.06† | 0.20‡ | 0.01 | 0.32‡ | 0.14‡ | 0.13‡ | 0.34‡ | 0.04 | 0 |
| South Africa | 0.01 | 0.09‡ | 0.02 | 0.16‡ | 0.03 | 0.04 | 0.27‡ | 0.09‡ | 0.02* |
| Breastfeeding duration (correlation) | | | | | | | | | |
| Brazil | 0.09‡ | 0.09‡ | 0.04* | -0.01 | -0.05‡ | 0.02 | 0.02 | 0.02 | 0 |
| Guatemala | 0.13‡ | 0.06 | 0.03 | 0.13‡ | -0.11† | -0.02 | 0.05 | 0.07 | 0.06 |
| Philippines | 0.04* | 0.05† | 0.04* | -0.09‡ | -0.12‡ | -0.20‡ | -0.05* | -0.01 | 0.03 |
| South Africa | -0.01 | 0.04 | 0.01 | -0.12‡ | 0.04 | 0 | 0 | 0.04 | -0.04 |

(Table 2 continues on next page)

| | Maternal age (years) | Birthweight (g) | Gestational age (weeks) | Height at 2 years (Z) | Weight-for-height at age 2 years (Z) | Schooling (years) | Adult height (cm) | Systolic blood pressure (mm Hg) | Fasting glucose (mmol/L) |
|--------------------------------|----------------------|-----------------|-------------------------|-----------------------|--------------------------------------|-------------------|-------------------|---------------------------------|--------------------------|
| (Continued from previous page) | | | | | | | | | |
| Parity (correlation) | | | | | | | | | |
| Brazil | 0.53‡ | 0.07‡ | -0.04* | -0.16‡ | -0.09‡ | -0.22‡ | -0.06‡ | -0.01 | 0 |
| Guatemala | 0.67‡ | 0.21‡ | 0.01 | -0.04 | 0.03 | -0.10‡ | -0.01 | 0 | 0.02 |
| India | 0.50‡ | 0.08‡ | 0.07* | -0.21‡ | -0.11‡ | -0.19‡ | -0.12‡ | -0.03 | -0.05 |
| Philippines | 0.64‡ | 0.17‡ | 0.03 | -0.20‡ | -0.06† | -0.14‡ | -0.03 | -0.01 | 0.04 |
| South Africa | 0.68‡ | 0.08‡ | 0.06† | -0.06* | -0.02 | -0.09‡ | 0.04 | 0.10‡ | 0 |

All participants in India were married. The Brazilian, Indian, and South African studies were set in exclusively urban areas. The Guatemala study was set in an exclusively rural area. Participants in India and the Philippines were classified as Asian and in Guatemala as other. Breastfeeding duration was not available for India. NA=data not available. *p<0.05. †p<0.01. ‡p<0.001. p values for marital status are from a 2-sample t test; those for ethnic origin are from a 3 degrees of freedom F test.

Table 2: Associations between potential confounding and mediating variables and maternal age and offspring birthweight, gestational age at delivery, size at age 2 years, schooling and adult height, systolic blood pressure, and fasting glucose between cohorts

maternal wealth and schooling values were imputed with regression analysis of known values on other socioeconomic variables. Missing maternal height values were not imputed, and a dummy variable (0=not missing; 1=missing) to represent missing value was included in regression models. The Guatemala cohort is based on a randomised controlled trial of a protein and energy supplement for pregnant women and children;²⁶ it comprises children living in the trial villages who were born or were younger than 7 years of age at any time between 1969 and 1977. We tested for interactions between maternal age and intervention group in this cohort, but noted no consistent evidence of interactions. In Guatemala, the 2344 participants came from 768 families; in India, the 5395 came from 5313 families; there were no siblings in the other cohorts. We used linear mixed modelling to assess whether siblings affected the associations of outcomes with maternal age, and found that they made little difference (appendix p 9). We therefore present our findings without adjustment for sibships. All analyses were done with SPSS version 21 and Stata version 12.

Role of the funding source

The funders had no role in the study design or conduct; the management, analysis, or interpretation of the data; the preparation, review, or approval of the report; or the decision to submit the manuscript for publication. CO and CHDF had full access to all the data and take responsibility for the integrity of the data and the accuracy of the data analysis. CHDF had the responsibility to submit the report for publication.

Results

We obtained data for 22188 mothers from the five cohorts, enrolment into which took place at various times between 1969 and 1989. Data for maternal age and at least one outcome were available for 19403 offspring (87%). Table 1 shows the characteristics of the

participants. Mean (range) maternal age was 26 years (12–49), which was similar in all cohorts; 20–24 years was the most numerous category (appendix p 2).

44% of mothers in South Africa were married and more than 90% were married in the other cohorts (table 1). Mean maternal schooling ranged from 1.3 years (Guatemala) to 9.5 years (South Africa). 55% of mothers in Guatemala were of parity 4 or higher compared with about 35% in India and the Philippines and 16% in Brazil and South Africa. Associations between the confounding or mediating variables and maternal age and study outcomes varied among the cohorts (table 2). Older maternal age was associated with being married, having less schooling but greater wealth, longer breastfeeding duration, and higher parity. Maternal height was positively related to maternal age in the Philippines, and we noted an inverted U-shaped relation in the other cohorts, with shorter height in mothers younger than 19 years and 35 years or older (non-linear data not shown). Maternal married status, higher schooling and wealth, white ethnic origin (Brazil and South Africa), urbanicity (the Philippines), and taller height were associated with higher birthweight, child size, and schooling. Blood pressure was highest in black participants and lowest in Asians (Brazil and South Africa).

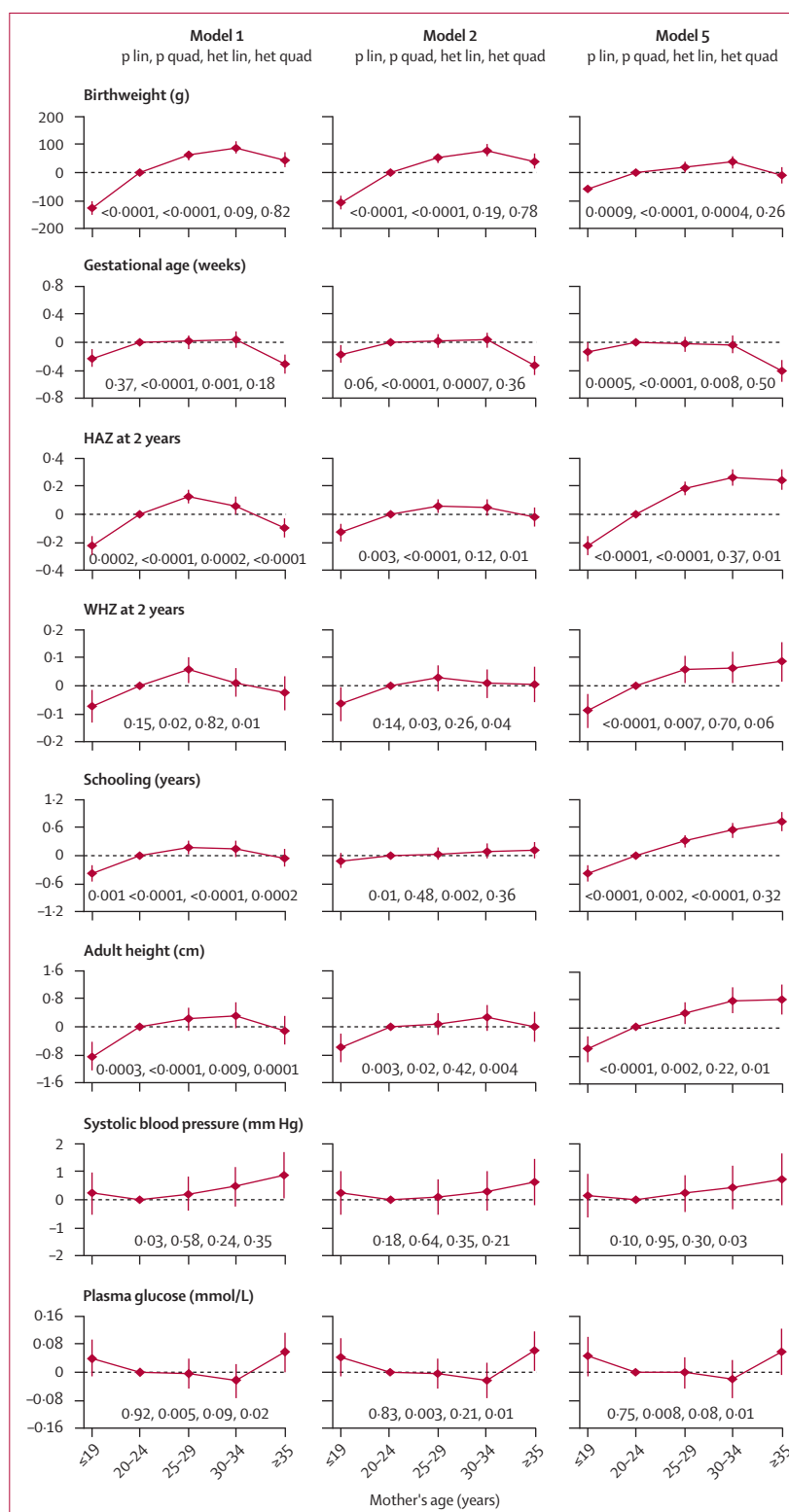
Birthweight rose with increasing maternal age, with a downturn among the oldest mothers (figure 1). Adjustment for socioeconomic factors, maternal height, and breastfeeding made little difference to this pattern. Adjustment for maternal parity attenuated the association, although maternal age 19 years or younger remained associated with lower birthweight (table 3). This association was similar for smallness-for-gestational-age (OR 1.10, 95% CI 0.94–1.24, in the fully adjusted model). Maternal age showed an inverted U-shaped association with gestational age at birth, which changed very little with adjustment (figure 1). Gestational age at birth was shorter by 0.1 weeks (95% CI 0.0–0.3)

in mothers aged 19 years and younger and 0·4 weeks (95% CI 0·3–0·6) in 35 years and older in the fully adjusted model. Odds ratios for preterm birth were 1·26 (95% CI 1·03–1·53) for mothers aged 19 years and younger and 1·33 (1·05–1·67) for mothers 35 years and older (table 3). These associations showed no site–sex heterogeneity (figure 1; appendix pp 5, 6, 10, 11).

Maternal age also showed inverted U-shaped associations with 2-year HAZ and WHZ, mid-childhood HAZ, attained schooling, and adult height (figure 1; appendix p 5). Adjustment for socioeconomic factors attenuated these associations, although maternal age 19 years or younger remained associated with lower values. Adjustment for maternal height and breastfeeding made little difference. Adjustment for parity changed the shape of the associations (figure 1); although young maternal age remained associated with poorer outcomes, maternal age 25 years or older was associated with better outcomes. In the fully adjusted model, 2-year HAZ was lower by 0·23 (95% CI 0·16–0·29) for maternal age 19 years or younger, and higher by 0·24 (95% CI 0·17–0·32) for maternal age 35 years or older. Corresponding data for schooling were –0·38 years (95% CI –0·54 to –0·21) and +0·72 years (95% CI 0·53 to 0·92) and for adult height –0·63 cm (–1·00 to –0·26) and +0·77 cm (0·34 to 1·20). The associations did not change much after further adjustment for birthweight. We noted similar findings for 2-year stunting and wasting, and failure to complete secondary schooling (table 3). Substantial site–sex heterogeneity was identified for these outcomes, mainly because of more linear associations in some cohorts, which were less marked in fully adjusted models (appendix pp 5, 6, 12–15).

Systolic blood pressure was positively and linearly related to maternal age (figure 1; appendix pp 5, 16), but this was non-significant after adjustment for socioeconomic factors. Fasting glucose showed a U-shaped association with maternal age (figure 1; appendix p 5), which remained after full adjustment. Mean fasting glucose was higher by 0·05 mmol/L (95% CI –0·01 to 0·10) in offspring of mothers aged 19 years and younger

and 0·06 mmol/L (–0·01 to 0·12) in offspring of mothers aged 35 years and older. This association did not change much after further adjustment for birthweight.



| | Low birthweight | Preterm birth | 2-year stunting | 2-year wasting | Non-completion of secondary schooling | Adult overweight or obesity | Adult high blood pressure | Adult impaired fasting glucose or diabetes |
|---|---------------------|---------------------|---------------------|---------------------|---------------------------------------|-----------------------------|---------------------------|--|
| Model 1—Maternal age (years; adjusted for sex only) | | | | | | | | |
| ≤19 | 1.43 (1.26–1.62) | 1.44 (1.21–1.72) | 1.37 (1.20–1.56) | 1.12 (0.86–1.46) | 1.32 (1.15–1.51) | 1.07 (0.92–1.25) | 1.00 (0.87–1.16) | 0.87 (0.64–1.18) |
| 20–24 | 1.00 (Ref) | 1.00 (Ref) | 1.00 (Ref) | 1.00 (Ref) | 1.00 (Ref) | 1.00 (Ref) | 1.00 (Ref) | 1.00 (Ref) |
| 25–29 | 0.83 (0.75–0.92) | 0.91 (0.78–1.06) | 0.84 (0.76–0.93) | 0.91 (0.75–1.11) | 0.80 (0.72–0.90) | 0.96 (0.84–1.08) | 1.05 (0.94–1.19) | 0.98 (0.78–1.23) |
| 30–34 | 0.82 (0.72–0.93) | 0.91 (0.77–1.08) | 0.93 (0.83–1.04) | 1.05 (0.84–1.31) | 0.86 (0.76–0.98) | 1.06 (0.92–1.21) | 1.11 (0.97–1.27) | 0.69 (0.52–0.91) |
| ≥35 | 0.94 (0.81–1.08) | 1.30 (1.08–1.57) | 1.18 (1.03–1.35) | 0.97 (0.74–1.28) | 0.99 (0.85–1.16) | 1.01 (0.86–1.19) | 1.08 (0.92–1.26) | 1.06 (0.79–1.41) |
| p* | <0.0001 | 0.29 | 0.02 | 0.64 | <0.0001 | 0.76 | 0.17 | 0.86 |
| p† | <0.0001 | <0.0001 | <0.0001 | 0.16 | <0.0001 | 0.74 | 0.35 | 0.29 |
| Model 2 (adjusted for sex and socioeconomic factors) | | | | | | | | |
| ≤19 | 1.36 (1.19–1.56) | 1.35 (1.12–1.62) | 1.23 (1.06–1.41) | 1.05 (0.80–1.38) | 1.13 (0.97–1.31) | 1.11 (0.95–1.29) | 1.00 (0.86–1.16) | 0.87 (0.64–1.19) |
| 20–24 | 1.00 (Ref) | 1.00 (Ref) | 1.00 (Ref) | 1.00 (Ref) | 1.00 (Ref) | 1.00 (Ref) | 1.00 (Ref) | 1.00 (Ref) |
| 25–29 | 0.87 (0.78–0.97) | 0.91 (0.77–1.06) | 0.88 (0.79–0.98) | 0.95 (0.77–1.16) | 0.87 (0.77–0.99) | 0.95 (0.84–1.08) | 1.04 (0.93–1.18) | 0.98 (0.78–1.23) |
| 30–34 | 0.83 (0.73–0.94) | 0.92 (0.77–1.10) | 0.87 (0.77–0.99) | 0.98 (0.78–1.23) | 0.87 (0.76–1.01) | 1.05 (0.91–1.21) | 1.10 (0.96–1.25) | 0.69 (0.52–0.92) |
| ≥35 | 0.92 (0.79–1.07) | 1.28 (1.04–1.56) | 0.97 (0.83–1.13) | 0.79 (0.59–1.06) | 0.88 (0.74–1.05) | 1.02 (0.87–1.21) | 1.06 (0.90–1.24) | 1.05 (0.78–1.43) |
| p* | <0.0001 | 0.69 | 0.0009 | 0.19 | 0.001 | 0.59 | 0.30 | 0.88 |
| p† | <0.0001 | <0.0001 | 0.003 | 0.92 | 0.10 | 0.56 | 0.35 | 0.22 |
| Model 5 (adjusted for sex, socioeconomic factors, maternal height, breastfeeding duration, and parity) | | | | | | | | |
| ≤19 | 1.18 (1.02–1.36) | 1.26 (1.03–1.53) | 1.46 (1.25–1.70) | 1.11 (0.83–1.49) | 1.38 (1.18–1.62) | 1.10 (0.94–1.29) | 0.95 (0.81–1.11) | 0.87 (0.63–1.20) |
| 20–24 | 1.00 (Ref) | 1.00 (Ref) | 1.009 (Ref) | 1.00 (Ref) | 1.00 (Ref) | 1.00 (Ref) | 1.00 (Ref) | 1.00 (Ref) |
| 25–29 | 0.96 (0.85–1.08) | 0.96 (0.81–1.14) | 0.72 (0.63–0.81) | 0.84 (0.68–1.04) | 0.71 (0.62–0.81) | 0.98 (0.86–1.12) | 1.07 (0.94–1.21) | 0.98 (0.77–1.25) |
| 30–34 | 0.94 (0.82–1.09) | 0.96 (0.79–1.18) | 0.63 (0.54–0.73) | 0.82 (0.63–1.06) | 0.63 (0.53–0.74) | 1.12 (0.96–1.30) | 1.13 (0.97–1.31) | 0.69 (0.51–0.94) |
| ≥35 | 1.03 (0.86–1.23) | 1.33 (1.05–1.67) | 0.64 (0.54–0.77) | 0.66 (0.48–0.90) | 0.59 (0.48–0.71) | 1.13 (0.94–1.36) | 1.08 (0.90–1.29) | 1.03 (0.74–1.44) |
| p* | 0.14 | 0.37 | <0.0001 | 0.01 | <0.0001 | 0.44 | 0.18 | 0.94 |
| p† | 0.001 | 0.0005 | <0.0001 | 0.32 | 0.0008 | 0.69 | 0.09 | 0.16 |

Data are odds ratio (95% CI). p*=p value for linear association. p†=p value for quadratic association. p values derived with logistic regression, with maternal age and maternal age² as continuous variables and outcomes as binary.

Table 3: Association between maternal age and low birthweight, preterm birth, 2-year stunting and wasting, non-completion of secondary education, adult overweight or obesity, high blood pressure, and abnormal glucose tolerance

Substantial site–sex heterogeneity was noted, due to more linear associations in some cohorts (appendix p 17). None of the body composition or other cardiometabolic outcomes was related to maternal age (table 3; appendix pp 5, 6).

The change in associations between maternal age and outcomes after adjustment for parity suggested strong effects of parity itself. Figure 2 shows these effects on birthweight, 2-year HAZ, and schooling. At any parity, birthweight increased with maternal age up to a threshold around 25–30 years and then decreased, whereas 2-year HAZ and schooling increased up to around 30–35 years, and remained high. At any maternal age, higher parity

was associated with higher birthweight (especially for the difference between parity 1 and 2), but lower 2-year HAZ and years of schooling. Higher parity was not associated with lower gestational age or higher risk of preterm birth in young or older mothers.

Discussion

With data from five population-based cohorts in LMICs, this study, to our knowledge, is one of the largest relating maternal age to birth and childhood outcomes with adjustment for multiple confounders, and the first to relate maternal age to adult outcomes (panel). Children of young mothers and mothers of advanced age had an

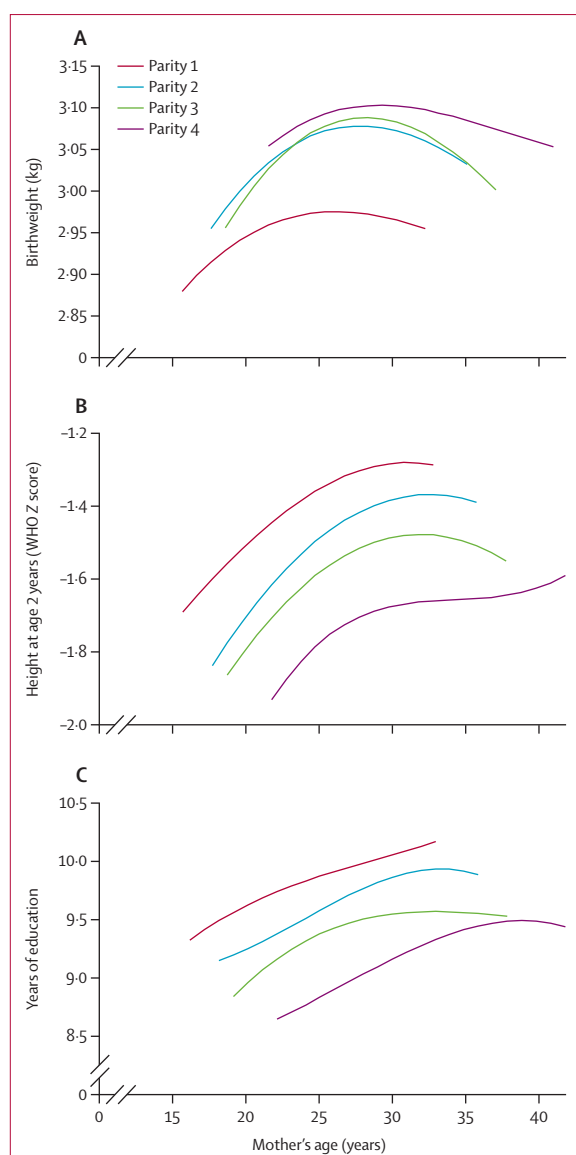


Figure 2: Relation between maternal age and parity and birthweight, 2-year height Z score, and years of schooling (adjusted for maternal socioeconomic status and height)

Outcomes (birthweight, height Z score at age 2 years, and schooling) were adjusted for study site and sex. Within each parity group (1, 2, 3, and 4+), cubic regression was used to model how maternal age was associated with each outcome. The fitted regression lines were plotted for the central 95% of maternal ages for the parity group.

increased risk of low birthweight and preterm birth, stunting in infancy, short adult height, poor schooling, and higher adult fasting glucose concentrations. The disadvantages of young maternal age were independent of socioeconomic status, height, breastfeeding duration, and parity. With the exception of preterm birth and adult plasma glucose concentration, the disadvantages associated with older maternal age were greatly reduced after adjustment for covariables, especially parity, and older maternal age was associated with distinct

Panel: Research in context

Systematic review

We searched PubMed for studies describing the relation between maternal age and birth outcomes, child nutritional status, attained schooling, and adult body composition and cardiometabolic risk factors in the offspring, published between Jan 1, 1990, and Jan 1, 2014, in English language. There is a large literature documenting an increased risk of adverse birth outcomes when the mother is very young (<19 years) or of advanced age (>35 years). Young motherhood is also associated with increased undernutrition and poor educational attainment in the children. This literature, especially that relating to advanced maternal age, is overwhelmingly dominated by studies from high-income countries. Few studies in any setting have examined advanced maternal age as a predictor of childhood outcomes, and none have examined young or advanced maternal age as predictors of adult body composition and cardiometabolic risk factors.

Interpretation

Our study pooled data from five birth cohorts in low-income and middle-income countries (LMICs) creating a large sample size (roughly 17 000 for birth outcomes, about 13 000 for child outcomes, and about 10 000 for adult outcomes). All studies collected data for maternal age at childbirth (exposure), and longitudinal measurements of the children's weight and height, attained schooling, and adult body composition and cardiometabolic risk factors (outcomes). They also collected data for maternal socioeconomic status, parity, height, and duration of breastfeeding, which might confound the relation of maternal age with these outcomes. We identified that young maternal age was associated with an increased risk of low birthweight, preterm and smallness-for-gestational-age births, stunting and wasting in infancy and childhood, and failure to complete secondary education. The associations were attenuated by adjustment for confounding factors, but remained significant, which suggests that they result from biological or behavioural immaturity of the mother. The results reinforce the arguments for LMICs to strengthen measures to prevent early marriage and childbearing. We confirmed that advanced maternal age was associated with an increased risk of preterm birth, independent of confounders. After adjustment for confounders, older maternal age was associated with improved child nutritional status and further schooling, suggesting that, other factors being equal, children of more mature mothers are advantaged nutritionally and educationally. Adult body composition and blood pressure were unrelated to maternal age, but fasting plasma glucose concentrations were increased in offspring of young and old mothers. This finding has not been reported before and should be replicated in other studies, but suggests that there could be effects of maternal age on fetal and child development, which programme adult glucose metabolism.

advantages for children's height and schooling. Apart from glucose, maternal age was unrelated to offspring cardiometabolic risk factors or body composition.

Children of mothers aged 19 years and younger had a 20–30% increased risk of low birthweight and preterm birth. This finding is consistent with a large body of evidence showing that teenage motherhood is detrimental to the newborn.^{1–12,36} Studies that subcategorised teenage births showed the most adverse effects at the youngest ages.^{2–4,6,9,11} The mechanisms for these effects are important for policy;⁸—for example, if they are explained by socioeconomic factors, interventions that simply delay age at marriage might have little effect. In our study, adjustment for socioeconomic factors attenuated the associations, but they remained significant. This suggests that there are biological or behavioural or a combination of both factors that are not captured by maternal height,

breastfeeding duration, or parity. This finding is supported by those noted in different settings.^{1–12} Several recent studies in the UK and the USA have reported that the increased risk of preterm birth among teenage mothers is greater for second-time than for first-time mothers.^{37,38} The reasons for this are unknown, but in a high-income setting, having more than one birth as a teenager could identify a group with high-risk behaviours. We did not note evidence of a similar association in our data from LMICs.

Children of mothers younger than 19 years had a 30–40% increased risk of 2-year stunting and of failure to complete secondary schooling. These associations remained after full adjustment, although residual confounding cannot be ruled out. The associations with poorer child nutritional status, especially stunting, are consistent with other studies.³⁹ The effects could be mediated by maternal inexperience, absence of autonomy leading to suboptimum feeding, and hygiene and health-care seeking behaviours,³ which we did not measure. Children of young mothers in LMICs have increased mortality, sometimes increasing with the age of the child, suggesting effects of the post natal environment.^{3,9} Studies from HICs have also shown poorer school attainment, cognitive function, or both in children of teenage mothers, independent of socioeconomic factors,^{20,40–42} possibly because of poorer care and stimulation. No previous publications have assessed schooling in relation to maternal age from LMICs.

In recognition of the disadvantages of young motherhood, most countries have a minimum legal age for marriage, but it is often ignored. In many LMICs, 20–30% of girls aged 15–19 years are married, and in some 40–50%.^{36,40} Early marriage has implications for women themselves, truncating their education and violating their human rights.^{36,43} The situation is improving in many places, but worsening in others, and the disparity in teenage fertility rates between HICs and LMICs is widening.²⁵ Our findings emphasise the importance of continued and strengthened measures to discourage early marriage and childbirth.

Independent of confounders, and consistent with other data from both HICs and LMICs,^{16–18} mothers 35 years or older had a 30% increased risk of preterm birth. However, in our fully adjusted data, older maternal age was an advantage for the child's nutritional status, schooling, and adult height. These associations were strong (the fully adjusted OR for non-completion of secondary education was 0·59) and have not been reported before from LMICs. They could be related to improved child-rearing practices by more experienced or empowered women.^{40–42}

We noted a U-shaped association between maternal age and fasting glucose concentration in the adult offspring. Offspring of younger and older mothers had concentrations of roughly 0·05 mmol/L higher than the reference group, an association that was not evident for impaired glucose tolerance or diabetes. This previously unreported association requires replication. The effect in

older mothers could be due to obesity, diabetes, or both in pregnancy, but we have no data to examine this. Offspring body composition and other cardiometabolic risk markers were unrelated to maternal age.

Strengths of the study are a large sample size, information about confounding factors recorded prospectively, and comparable outcomes in all sites. We did not know maternal age at menarche and so could not examine effects of gynaecological age. We did not have data for maternal smoking, diet, and supplement use, other aspects of infant or childhood diet (apart from breastfeeding), and paternal factors, and so residual confounding cannot be ruled out. We cannot exclude confounding by unmeasured genetic or familial factors, which have been recently shown in a large Swedish cohort using sibling comparisons and cousin comparisons to account for an association between young maternal age and attention-deficit hyperactivity disorder in the children.⁴⁴ Adult outcomes were measured with different techniques in the different cohorts, although all were accepted methods; these differences could contribute additional noise in the data, which would tend to reduce associations. Other limitations were losses to follow-up and missing data. 2-year outcomes were available for nearly 80% of the cohort in Brazil and the Philippines, but only 45–50% in Guatemala (where many children were enrolled after age 2 years), India, and South Africa. Percentages for adult blood pressure were greater than 70% in Brazil, 50% in Guatemala, the Philippines, and South Africa, and 14% in India. Comparisons between participants and non-participants showed differences among cohorts in maternal, child, and adult measurements that varied in strength and direction; we adjusted for each maternal factor independently within each cohort to allow for the varied ways in which they were related to outcomes.

Children of teenage mothers in LMICs are disadvantaged at birth and during childhood and have reduced human capital. Measures to prevent young motherhood should be strengthened and young mothers should be helped to improve their children's nutrition and education. Our data show that children of older mothers are at increased risk of preterm birth, and might have increased plasma glucose concentrations in adult life. However, at a given socioeconomic level and parity, children of older mothers have advantages in terms of childhood nutrition and educational attainment.

Contributors

The COHORTS group designed the research. CHDF, HSS, CO, MCR-M, CV, RM, ADS, SS, NT, LA, IB, SN, and LMR collected the data. CO and CHDF analysed the data and drafted the report. All authors contributed to critical revision of the report. CO and CHDF had primary responsibility for final content. All authors read and approved the final report.

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Declaration of interests

We declare no competing interests.

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